Solow Growth Model and the Data

- Use Solow model or extensions to interpret both economic growth over time and cross-country output differences.
- Focus on *proximate causes* of economic growth.
Growth Accounting I

- Aggregate production function in its general form:
  \[ Y(t) = F[K(t), L(t), A(t)]. \]
- Continuous-time economy and differentiate the aggregate production function with respect to time.
- Dropping time dependence,
  \[
  \frac{\dot{Y}}{Y} = \frac{F_A A}{A} \frac{\dot{A}}{A} + \frac{F_K K}{K} \frac{\dot{K}}{K} + \frac{F_L L}{L} \frac{\dot{L}}{L}. \tag{1}
  \]
Growth Accounting II

- Denote growth rates of output, capital stock and labor by $g \equiv \dot{Y}/Y$, $g_K \equiv \dot{K}/K$ and $g_L \equiv \dot{L}/L$.
- Define the contribution of technology to growth as
  \[ x \equiv \frac{FAA \dot{A}}{Y A} \]
- Recall with competitive factor markets, $w = F_L$ and $R = F_K$.
- Define factor shares as $\alpha_K \equiv RK/Y$ and $\alpha_L \equiv wL/Y$.
- Putting all these together, (1) the fundamental growth accounting equation
  \[ x = g - \alpha_K g_K - \alpha_L g_L. \]  
- Gives estimate of contribution of technological progress, Total Factor Productivity (TFP) or Multi Factor Productivity as
  \[ \hat{x}(t) = g(t) - \alpha_K(t) g_K(t) - \alpha_L(t) g_L(t). \]  
- All terms on right-hand side are “estimates” obtained with a range of assumptions from national accounts and other data sources.
Growth Accounting III

- In continuous time, equation (3) is exact.
- With discrete time, potential problem in using (3): over the time horizon factor shares can change.
- Use beginning-of-period or end-of-period values of $\alpha_K$ and $\alpha_L$?
  - Either might lead to seriously biased estimates.
  - Best way of avoiding such biases is to use as high-frequency data as possible.
  - Typically use factor shares calculated as the average of the beginning and end of period values.

- In discrete time, the analog of equation (3) becomes

$$\hat{x}_{t,t+1} = g_{t,t+1} - \bar{\alpha}_K,t,t+1 g_K,t,t+1 - \bar{\alpha}_L,t,t+1 g_L,t,t+1, \quad (4)$$

- $g_{t,t+1}$ is the growth rate of output between $t$ and $t + 1$; other growth rates defined analogously.
Moreover,

\[
\bar{\alpha}_{K,t,t+1} \equiv \frac{\alpha_K(t) + \alpha_K(t+1)}{2}
\]

and

\[
\bar{\alpha}_{L,t,t+1} \equiv \frac{\alpha_L(t) + \alpha_L(t+1)}{2}
\]

Equation (4) would be a fairly good approximation to (3) when the difference between \( t \) and \( t + 1 \) is small and the capital-labor ratio does not change much during this time interval.

Solow's (1957) applied this framework to US data: a large part of the growth was due to technological progress.

From early days, however, a number of pitfalls were recognized.

- Moses Abramovitz (1956): dubbed the \( \hat{x} \) term "the measure of our ignorance".
- If we mismeasure \( g_L \) and \( g_K \) we will arrive at inflated estimates of \( \hat{x} \).
Growth Accounting Results

- Example from Barro and Sala-i-Martin’s textbook

### Table 10.1 Growth Accounting for a Sample of Countries

<table>
<thead>
<tr>
<th>Country</th>
<th>(1) Growth Rate of GDP</th>
<th>(2) Contribution from Capital</th>
<th>(3) Contribution from Labor</th>
<th>(4) TFP Growth Rate</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Panel A: OECD Countries, 1947-73</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Canada</td>
<td>0.0517</td>
<td>0.0254</td>
<td>0.0088</td>
<td>0.0175</td>
</tr>
<tr>
<td>(α = 0.44)</td>
<td></td>
<td>(49%)</td>
<td>(17%)</td>
<td>(34%)</td>
</tr>
<tr>
<td>France</td>
<td>0.0542</td>
<td>0.0225</td>
<td>0.0021</td>
<td>0.0296</td>
</tr>
<tr>
<td>(α = 0.40)</td>
<td></td>
<td>(42%)</td>
<td>(4%)</td>
<td>(54%)</td>
</tr>
<tr>
<td>Germany</td>
<td>0.0661</td>
<td>0.0269</td>
<td>0.0018</td>
<td>0.0374</td>
</tr>
<tr>
<td>(α = 0.39)</td>
<td></td>
<td>(41%)</td>
<td>(4%)</td>
<td>(56%)</td>
</tr>
<tr>
<td>Italy</td>
<td>0.0527</td>
<td>0.0180</td>
<td>0.0011</td>
<td>0.0337</td>
</tr>
<tr>
<td>(α = 0.39)</td>
<td></td>
<td>(34%)</td>
<td>(3%)</td>
<td>(64%)</td>
</tr>
<tr>
<td>Japan</td>
<td>0.0951</td>
<td>0.0328</td>
<td>0.0221</td>
<td>0.0402</td>
</tr>
<tr>
<td>(α = 0.39)</td>
<td></td>
<td>(35%)</td>
<td>(23%)</td>
<td>(42%)</td>
</tr>
<tr>
<td>Netherlands</td>
<td>0.0536</td>
<td>0.0247</td>
<td>0.0042</td>
<td>0.0248</td>
</tr>
<tr>
<td>(α = 0.45)</td>
<td></td>
<td>(46%)</td>
<td>(8%)</td>
<td>(46%)</td>
</tr>
<tr>
<td>U.K.</td>
<td>0.0373</td>
<td>0.0176</td>
<td>0.0003</td>
<td>0.0193</td>
</tr>
<tr>
<td>(α = 0.38)</td>
<td></td>
<td>(47%)</td>
<td>(1%)</td>
<td>(52%)</td>
</tr>
<tr>
<td>U.S.</td>
<td>0.0402</td>
<td>0.0171</td>
<td>0.0095</td>
<td>0.0135</td>
</tr>
<tr>
<td>(α = 0.40)</td>
<td></td>
<td>(43%)</td>
<td>(24%)</td>
<td>(34%)</td>
</tr>
</tbody>
</table>

| **Panel B: OECD Countries, 1960-95** |                        |                               |                             |                     |
| Canada        | 0.0369                 | 0.0186                        | 0.0123                      | 0.0057              |
| (α = 0.42)    |                        | (51%)                         | (33%)                       | (16%)               |
| France        | 0.0358                 | 0.0180                        | 0.0033                      | 0.0130              |
| (α = 0.41)    |                        | (53%)                         | (10%)                       | (38%)               |
| Germany       | 0.0312                 | 0.0177                        | 0.0014                      | 0.0132              |
| (α = 0.39)    |                        | (56%)                         | (4%)                        | (42%)               |
| Italy         | 0.0357                 | 0.0182                        | 0.0035                      | 0.0153              |
| (α = 0.34)    |                        | (51%)                         | (9%)                        | (42%)               |
| Japan         | 0.0566                 | 0.0178                        | 0.0125                      | 0.0265              |
| (α = 0.43)    |                        | (51%)                         | (22%)                       | (47%)               |
| U.K.          | 0.0221                 | 0.0124                        | 0.0017                      | 0.0080              |
| (α = 0.37)    |                        | (56%)                         | (8%)                        | (36%)               |
| U.S.          | 0.0318                 | 0.0117                        | 0.0127                      | 0.0076              |
| (α = 0.39)    |                        | (37%)                         | (40%)                       | (24%)               |

*Table continued*
### Growth Accounting Results (continued)

#### Table 10.1 (Continued)

<table>
<thead>
<tr>
<th>Country</th>
<th>(1) Growth Rate of GDP</th>
<th>(2) Contribution from Capital</th>
<th>(3) Contribution from Labor</th>
<th>(4) TFP Growth Rate</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Panel C: Latin American Countries, 1940–90</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Argentina</td>
<td>0.0279</td>
<td>0.0128</td>
<td>0.0097</td>
<td>0.0054</td>
</tr>
<tr>
<td>(\alpha = 0.54)</td>
<td></td>
<td>0.46%</td>
<td>(35%)</td>
<td>(19%)</td>
</tr>
<tr>
<td>Brazil</td>
<td>0.0558</td>
<td>0.0294</td>
<td>0.0150</td>
<td>0.0114</td>
</tr>
<tr>
<td>(\alpha = 0.45)</td>
<td></td>
<td>0.53%</td>
<td>(27%)</td>
<td>(20%)</td>
</tr>
<tr>
<td>Chile</td>
<td>0.0362</td>
<td>0.0120</td>
<td>0.0103</td>
<td>0.0138</td>
</tr>
<tr>
<td>(\alpha = 0.52)</td>
<td></td>
<td>0.33%</td>
<td>(28%)</td>
<td>(38%)</td>
</tr>
<tr>
<td>Colombia</td>
<td>0.0454</td>
<td>0.0219</td>
<td>0.0152</td>
<td>0.0084</td>
</tr>
<tr>
<td>(\alpha = 0.63)</td>
<td></td>
<td>0.48%</td>
<td>(33%)</td>
<td>(19%)</td>
</tr>
<tr>
<td>Mexico</td>
<td>0.0522</td>
<td>0.0259</td>
<td>0.0150</td>
<td>0.0113</td>
</tr>
<tr>
<td>(\alpha = 0.69)</td>
<td></td>
<td>0.50%</td>
<td>(29%)</td>
<td>(22%)</td>
</tr>
<tr>
<td>Peru</td>
<td>0.0323</td>
<td>0.0252</td>
<td>0.0134</td>
<td>-0.0062</td>
</tr>
<tr>
<td>(\alpha = 0.66)</td>
<td></td>
<td>0.78%</td>
<td>(41%)</td>
<td>(−19%)</td>
</tr>
<tr>
<td>Venezuela</td>
<td>0.0443</td>
<td>0.0254</td>
<td>0.0179</td>
<td>0.0011</td>
</tr>
<tr>
<td>(\alpha = 0.55)</td>
<td></td>
<td>0.57%</td>
<td>(40%)</td>
<td>(2%)</td>
</tr>
<tr>
<td><strong>Panel D: East Asian Countries, 1966–90</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hong Kong(^a)</td>
<td>0.073</td>
<td>0.030</td>
<td>0.020</td>
<td>0.023</td>
</tr>
<tr>
<td>(\alpha = 0.37)</td>
<td></td>
<td>(41%)</td>
<td>(28%)</td>
<td>(32%)</td>
</tr>
<tr>
<td>Singapore</td>
<td>0.087</td>
<td>0.056</td>
<td>0.029</td>
<td>0.002</td>
</tr>
<tr>
<td>(\alpha = 0.49)</td>
<td></td>
<td>0.65%</td>
<td>(33%)</td>
<td>(2%)</td>
</tr>
<tr>
<td>South Korea</td>
<td>0.103</td>
<td>0.041</td>
<td>0.045</td>
<td>0.017</td>
</tr>
<tr>
<td>(\alpha = 0.30)</td>
<td></td>
<td>0.40%</td>
<td>(44%)</td>
<td>(16%)</td>
</tr>
<tr>
<td>Taiwan</td>
<td>0.094</td>
<td>0.032</td>
<td>0.036</td>
<td>0.026</td>
</tr>
<tr>
<td>(\alpha = 0.26)</td>
<td></td>
<td>0.34%</td>
<td>(39%)</td>
<td>(28%)</td>
</tr>
</tbody>
</table>

Source: Panel A estimates for OECB countries are from Clark et al. (2009) and Romer (2005).
Reasons for mismeasurement:

- what matters is not labor hours, but effective labor hours
  - important—though difficult—to make adjustments for changes in the human capital of workers.

measurement of capital inputs:

- in the theoretical model, capital corresponds to the final good used as input to produce more goods.
- in practice, capital is machinery, need assumptions about how relative prices of machinery change over time.
- typical assumption was to use capital expenditures but if machines become cheaper would severely underestimate $g_K$
Mankiw, Romer and Weil (1992) used regression analysis to take the augmented Solow model, with human capital, to data.

Use the Cobb-Douglas model and envisage a world consisting of \( j = 1, \ldots, N \) countries.

“Each country is an island”: countries do not interact (perhaps except for sharing some common technology growth).

Country \( j = 1, \ldots, N \) has the aggregate production function:

\[
Y_j(t) = K_j(t)^{\alpha} H_j(t)^{\beta} (A_j(t) L_j(t))^{1-\alpha-\beta}.
\]

Nests the basic Solow model without human capital when \( \alpha = 0 \).

Countries differ in terms of their saving rates, \( s_{k,j} \) and \( s_{h,j} \), population growth rates, \( n_j \), and technology growth rates \( \dot{A}_j(t)/A_j(t) = g_j \).

Define \( k_j \equiv K_j/A_jL_j \) and \( h_j \equiv H_j/A_jL_j \).
A World of Augmented Solow Economies II

- Focus on a world in which each country is in their steady state
- Assuming that human capital also has depreciation, at the rate $\delta_h$, and it is accumulated with the saving rate $s_h$, steady state values for country $j$ would be (to be derived in recitation):

$$k_j^* = \left( \left( \frac{s_{k,j}}{n_j + g_j + \delta_k} \right)^{1-\beta} \left( \frac{s_{h,j}}{n_j + g_j + \delta_h} \right)^{\beta} \right)^{\frac{1}{1-\alpha-\beta}}$$

$$h_j^* = \left( \left( \frac{s_{k,j}}{n_j + g_j + \delta_k} \right)^{\alpha} \left( \frac{s_{h,j}}{n_j + g_j + \delta_h} \right)^{1-\alpha} \right)^{\frac{1}{1-\alpha-\beta}}.$$

- Consequently:

$$y_j^* (t) \equiv \frac{Y(t)}{L(t)} = A_j(t) \left( \frac{s_{k,j}}{n_i + g_i + \delta_k} \right)^{\frac{\alpha}{1-\alpha-\beta}} \left( \frac{s_{h,i}}{n_i + g_i + \delta_h} \right)^{\frac{\beta}{1-\alpha-\beta}}.$$
Here $y_j^* (t)$ stands for output per capita of country $j$ along the balanced growth path.

Note if $g_j$’s are not equal across countries, income per capita will diverge.

Mankiw, Romer and Weil (1992) make the following assumption:

$$ A_j (t) = \bar{A}_j \exp (gt). $$

Countries differ according to technology level, (initial level $\bar{A}_j$) but they share the same common technology growth rate, $g$. 
Using this together with (5) and taking logs, equation for the balanced growth path of income for country \( j = 1, \ldots, N \):

\[
\ln y_j^* (t) = \ln \bar{A}_j + gt + \frac{\alpha}{1 - \alpha - \beta} \ln \left( \frac{s_{k,j}}{n_j + g + \delta_k} \right) + \frac{\beta}{1 - \alpha - \beta} \ln \left( \frac{s_{h,j}}{n_j + g + \delta_h} \right).
\]  

Mankiw, Romer and Weil (1992) take:

- \( \delta_k = \delta_h = \delta \) and \( \delta + g = 0.05 \).
- \( s_{k,j} \) = average investment rates (investments/GDP).
- \( s_{h,j} \) = fraction of the school-age population that is enrolled in secondary school.
Even with all of these assumptions, (6) can still not be estimated consistently.

In $\bar{A}_j$ is unobserved (at least to the econometrician) and thus will be captured by the error term.

Most reasonable models would suggest $\ln \bar{A}_j$’s should be correlated with investment rates.

Thus an estimation of (6) would lead to omitted variable bias and inconsistent estimates.

Implicitly, MRW make another crucial assumption, the orthogonal technology assumption:

$$\bar{A}_j = \varepsilon_j A,$$
with $\varepsilon_j$ orthogonal to all other variables.
MRW first estimate equation (6) without the human capital term for the cross-sectional sample of non-oil producing countries

\[
\ln y_j^* = \text{constant} + \frac{\alpha}{1 - \alpha} \ln (s_{k,j}) - \frac{\alpha}{1 - \alpha} \ln (n_j + g + \delta_k) + \varepsilon_j.
\]
## Cross-Country Income Differences: Regressions II

### Estimates of the Basic Solow Model

<table>
<thead>
<tr>
<th></th>
<th>MRW 1985</th>
<th>Updated data 1985</th>
<th>Updated data 2000</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \ln(s_k) )</td>
<td>1.42 (.14)</td>
<td>1.01 (.11)</td>
<td>1.22 (.13)</td>
</tr>
<tr>
<td>( \ln(n + g + \delta) )</td>
<td>-1.97 (.56)</td>
<td>-1.12 (.55)</td>
<td>-1.31 (.36)</td>
</tr>
<tr>
<td>Adj ( R^2 )</td>
<td>.59</td>
<td>.49</td>
<td>.49</td>
</tr>
<tr>
<td>Implied ( \alpha )</td>
<td>.59</td>
<td>.50</td>
<td>.55</td>
</tr>
<tr>
<td>No. of observations</td>
<td>98</td>
<td>98</td>
<td>107</td>
</tr>
</tbody>
</table>
Cross-Country Income Differences: Regressions III

- Their estimates for \( \alpha / (1 - \alpha) \), implies that \( \alpha \) must be around 2/3, but should be around 1/3.

- The most natural reason for the high implied values of \( \alpha \) is that \( \varepsilon_j \) is correlated with \( \ln(s_{k,j}) \), either because:
  1. the orthogonal technology assumption is not a good approximation to reality or
  2. there are also human capital differences correlated with \( \ln(s_{k,j}) \).

- Mankiw, Romer and Weil favor the second interpretation and estimate the augmented model,

\[
\ln y_j^* = \text{cst} + \frac{\alpha}{1 - \alpha - \beta} \ln (s_{k,j}) - \frac{\alpha}{1 - \alpha - \beta} \ln (n_j + g + \delta_k) \tag{7} \\
+ \frac{\beta}{1 - \alpha - \beta} \ln (s_{h,j}) - \frac{\beta}{1 - \alpha - \beta} \ln (n_j + g + \delta_h) + \varepsilon_j.
\]
### Estimates of the Augmented Solow Model

<table>
<thead>
<tr>
<th></th>
<th>MRW 1985</th>
<th>Updated data 1985</th>
<th>2000</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\ln(s_k)$</td>
<td>.69</td>
<td>.65</td>
<td>.96</td>
</tr>
<tr>
<td>($\text{SD}$)</td>
<td>(.13)</td>
<td>(.11)</td>
<td>(.13)</td>
</tr>
<tr>
<td>$\ln(n + g + \delta)$</td>
<td>-1.73</td>
<td>-1.02</td>
<td>-1.06</td>
</tr>
<tr>
<td>($\text{SD}$)</td>
<td>(.41)</td>
<td>(.45)</td>
<td>(.33)</td>
</tr>
<tr>
<td>$\ln(s_h)$</td>
<td>.66</td>
<td>.47</td>
<td>.70</td>
</tr>
<tr>
<td>($\text{SD}$)</td>
<td>(.07)</td>
<td>(.07)</td>
<td>(.13)</td>
</tr>
<tr>
<td>$\text{Adj R}^2$</td>
<td>.78</td>
<td>.65</td>
<td>.60</td>
</tr>
<tr>
<td>$\text{Implied } \alpha$</td>
<td>.30</td>
<td>.31</td>
<td>.36</td>
</tr>
<tr>
<td>$\text{Implied } \beta$</td>
<td>.28</td>
<td>.22</td>
<td>.26</td>
</tr>
<tr>
<td>$\text{No. of observations}$</td>
<td>98</td>
<td>98</td>
<td>107</td>
</tr>
</tbody>
</table>
If these regression results are reliable, they give a big boost to the augmented Solow model.

- Adjusted $R^2$ suggests that three quarters of income per capita differences across countries can be explained by differences in their physical and human capital investment.

- Immediate implication is technology (TFP) differences have a somewhat limited role.

- But this conclusion should not be accepted without further investigation.
Challenges to Regression Analyses I

- **Technology differences across countries are not orthogonal to all other variables.**
- \( \bar{A}_j \) is correlated with measures of \( s^h_j \) and \( s^k_j \) for two reasons.
  1. **omitted variable bias**: societies with high \( \bar{A}_j \) will be those that have invested more in technology for various reasons; same reasons likely to induce greater investment in physical and human capital as well.
  2. **reverse causality**: complementarity between technology and physical or human capital imply that countries with high \( \bar{A}_j \) will find it more beneficial to increase their stock of human and physical capital.
- In terms of \((7)\), implies that key right-hand side variables are correlated with the error term, \( \varepsilon_j \).
- OLS estimates of \( \alpha \) and \( \beta \) and \( R^2 \) are biased upwards.
Challenges to Regression Analyses II

- $\beta$ is too large relative to what we should expect on the basis of microeconometric evidence.
- The working age population enrolled in school ranges from 0.4% to over 12% in the sample of countries.
- Predicted log difference in incomes between these two countries is

$$\frac{\beta}{1 - \alpha - \beta} \left( \ln 12 - \ln (0.4) \right) = 0.66 \times \left( \ln 12 - \ln (0.4) \right) \approx 2.24.$$  

- Thus a country with schooling investment of over 12 should be about $\exp(2.24) - 1 \approx 8.5$ times richer than one with investment of around 0.4.
Take Mincer regressions of the form:

\[ \ln w_i = X_i' \gamma + \phi S_i, \]  

(8)

Microeconometrics literature suggests that \( \phi \) is between 0.06 and 0.10.

Can deduce how much richer a country with 12 if we assume:

1. That the micro-level relationship as captured by (8) applies identically to all countries.
2. That there are no human capital externalities.

Then: a country with 12 more years of average schooling should have between \( \exp(0.10 \times 12) \approx 3.3 \) and \( \exp(0.06 \times 12) \approx 2.05 \) times the stock of human capital of a county with fewer years of schooling.
Thus holding other factors constant, this country should be about 2-3 times as rich as the country with zero years of average schooling.

Much less than the 8.5 fold difference implied by the Mankiw-Romer-Weil analysis.

Thus $\beta$ in MRW is too high relative to the estimates implied by the microeconometric evidence and thus likely upwardly biased.

Overestimation of $\beta$ is, in turn, most likely related to correlation between the error term $\varepsilon_j$ and the key right-hand side regressors in (7).

Return to basic Solow model with constant population growth and labor-augmenting technological change in continuous time:

\[ y(t) = A(t) f(k(t)), \quad (9) \]

and

\[ \frac{\dot{k}(t)}{k(t)} = \frac{sf(k(t))}{k(t)} - \delta - g - n, \quad (10) \]
Differentiating (9) with respect to time and dividing both sides by $y(t)$,

$$
\frac{\dot{y}(t)}{y(t)} = g + \varepsilon_f(k(t)) \frac{\dot{k}(t)}{k(t)},
$$

where

$$
\varepsilon_f(k(t)) \equiv \frac{f'(k(t)) k(t)}{f(k(t))} \in (0, 1)
$$

is the elasticity of the $f(\cdot)$ function.

$\varepsilon_f(k(t))$ is between 0 and 1 follows from Assumption 1. For example, with Cobb-Douglas $\varepsilon_f(k(t)) = \alpha$, but generally a function of $k(t)$. 
First-order Taylor expansion of (10) with respect to \( \log k(t) \) around \( k^* \) (and recall that \( \frac{\partial y}{\partial \log x} = (\frac{\partial y}{\partial x}) \cdot x \)):

\[
\frac{\dot{k}(t)}{k(t)} \simeq \left( \frac{sf(k^*)}{k^*} - \delta - g - n \right) + \left( \frac{f'(k^*) k^*}{f(k^*)} - 1 \right) s \frac{f(k^*)}{k^*} \left( \log k(t) - \log k^* \right).
\]

\[
\simeq (\varepsilon_f(k^*) - 1) (\delta + g + n) \left( \log k(t) - \log k^* \right).
\]

First term in the first line is zero by definition of the steady-state value \( k^* \).

Also used definition of \( \varepsilon_f(k(t)) \) and the fact that \( sf(k^*) / k^* = \delta + g + n \).

Substituting into (11),

\[
\frac{\dot{y}(t)}{y(t)} \simeq g - \varepsilon_f(k^*) (1 - \varepsilon_f(k^*)) (\delta + g + n) \left( \log k(t) - \log k^* \right).
\]
Solow Model and Growth Regressions IV

- Define $y^* (t) \equiv A(t) f(k^*)$; refer to $y^* (t)$ as the “steady-state level of output per capita” even though it is not constant.
- First-order Taylor expansions of $\log y(t)$ with respect to $\log k(t)$ around $\log k^*(t)$:

  $$\log y(t) - \log y^*(t) \simeq \varepsilon_f (k^*) (\log k(t) - \log k^*).$$

- Combining this with the previous equation, “convergence equation”:

  $$\frac{\dot{y}(t)}{y(t)} \simeq g - (1 - \varepsilon_f (k^*)) (\delta + g + n) (\log y(t) - \log y^*(t)). \quad (12)$$

- Two sources of growth in Solow model: $g$, the rate of technological progress, and “convergence”.

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**Daron Acemoglu (MIT) Economic Growth Lecture 4 November 8, 2016. 27 / 42**
Latter source, convergence:

- Negative impact of the gap between current level and steady-state level of output per capita on rate of capital accumulation (recall $0 < \varepsilon_f(k^*) < 1$).
- The lower is $y(t)$ relative to $y^*(t)$, hence the lower is $k(t)$ relative to $k^*$, the greater is $f(k^*)/k^*$, and this leads to faster growth in the effective capital-labor ratio.

Speed of convergence in (12), measured by the term 
\[(1 - \varepsilon_f(k^*)) (\delta + g + n),\]
depends on:

- $\delta + g + n$: determines rate at which effective capital-labor ratio needs to be replenished.
- $\varepsilon_f(k^*)$: when $\varepsilon_f(k^*)$ is high, we are close to a linear—$AK$—production function, convergence should be slow.
Example: Cobb-Douglas Production Function

- Consider Cobb-Douglas production function
  \[ Y(t) = A(t) K(t)^\alpha L(t)^{1-\alpha}. \]

- Then (12) becomes
  \[ \frac{\dot{y}(t)}{y(t)} \simeq g - (1 - \alpha) (\delta + g + n) (\log y(t) - \log y^*(t)). \]

- Focus on advanced economies for a back of the envelope calculation:
  - \( g \simeq 0.02 \) for approximately 2% per year output per capita growth,
  - \( n \simeq 0.01 \) for approximately 1% population growth and
  - \( \delta \simeq 0.05 \) for about 5% per year depreciation.

  - Share of capital in national income is about 1/3, so \( \alpha \simeq 1/3. \)

- Thus convergence coefficient would be around 0.054 (\( \simeq 0.67 \times 0.08 \)), which is very rapid relative to what some authors estimate from cross-country regressions.
Solow Model and Growth Regressions VI

- Using (12), we can obtain a growth regression similar to those estimated by Barro (1991).
- Using discrete time approximations, equation (12) yields:

\[ g_{i,t,t-1} = b^0 + b^1 \log y_{i,t-1} + \varepsilon_{i,t}, \]  

(13)

where \( \varepsilon_{i,t} \) is a stochastic term capturing all omitted influences.

- If such an equation is estimated in the sample of core OECD countries, \( b^1 \) is indeed estimated to be negative. But for the whole world, no evidence for a negative \( b^1 \). If anything, \( b^1 \) would be positive, i.e., there is no evidence of world-wide convergence,

- Barro and Sala-i-Martin refer to this as “unconditional convergence.” But this might be too demanding:

  - requires income gap between two countries to decline, irrespective of what types of technological opportunities, policies and institutions these countries have. If countries do differ, Solow model would not predict that they should converge in income level.
If countries differ according to characteristics, then perhaps
\[ g_{i,t-1} = b^0_i + b^1 \log y_{i,t-1} + \varepsilon_{i,t}, \]  
(14)

Now the constant term, \( b^0_i \), is country specific, and can be, for example, modeled as
\[ b^0_i = X'_{i,t} \beta + \delta_i + u_{i,t}, \]
where \( \delta_i \) denotes country fixed effects.

In this case, focus is on “conditional convergence,” i.e., on whether \( b^1 < 0 \).

This equation can be estimated using panel data methods as in the first lecture, but **much care is necessary**.

\( X_{i,t} \) should not include channels (such as education and investment); lots of biases and causality definitely not guaranteed. If these problems exist for the model is not specified properly, \( b^1 \) will not be estimated consistently.
Suppose each country has access to the Cobb-Douglas aggregate production function:

\[ Y_j = K_j^\alpha (A_j H_j)^{1-\alpha} \]  

(15)

Each worker in country \( j \) has \( S_j \) years of schooling.

Then using the Mincer equation (8) ignoring the other covariates and taking exponents, \( H_j \) can be estimated as

\[ H_j = \exp (\phi S_j) L_j, \]

Does not take into account differences in other “human capital” factors, such as experience.
Calibrating Productivity Differences II

- Let the rate of return to acquiring the $S$th year of schooling be $\phi(S)$.
- A better estimate of the stock of human capital can be constructed as

$$H_j = \sum_S \exp \{ \phi(S) S \} L_j(S)$$

- $L_j(S)$ now refers to the total employment of workers with $S$ years of schooling in country $j$.
- Series for $K_j$ can be constructed from Summers-Heston dataset using investment data and the perpetual inventory method.

$$K_j(t + 1) = (1 - \delta) K_j(t) + I_j(t),$$

- Assume, following Hall and Jones that $\delta = 0.06$.
- With same arguments as before, choose a value of $1/3$ for $\alpha$. 

Calibrating Productivity Differences III

- Given series for $H_j$ and $K_j$ and a value for $\alpha$, construct “predicted” incomes at a point in time using

$$\hat{Y}_j = K_j^{1/3} (A_{US} H_j)^{2/3}$$

- $A_{US}$ is computed so that $Y_{US} = K_{US}^{1/3} (A_{US} H_{US})^{2/3}$.

- Once a series for $\hat{Y}_j$ has been constructed, it can be compared to the actual output series.

- Gap between the two series represents the contribution of technology.

- Alternatively, could back out country-specific technology terms (relative to the United States) as

$$\frac{A_j}{A_{US}} = \left( \frac{Y_j}{Y_{US}} \right)^{3/2} \left( \frac{K_{US}}{K_j} \right)^{1/2} \left( \frac{H_{US}}{H_j} \right).$$
Calibrating Productivity Differences IV

Figure: Calibrated technology levels relative to the US technology (from the Solow growth model with human capital) versus log GDP per worker, 1980, 1990 and 2000.
Figure: Calibrated technology levels relative to the US technology (from the Solow growth model with human capital) versus log GDP per worker, 1980, 1990 and 2000.
The following features are noteworthy:

1. Differences in physical and human capital still matter a lot.

2. However, differently from the regression analysis, this exercise also shows significant *technology (productivity) differences*.

3. Same pattern visible in the next three figures for the estimates of the technology differences, $A_j/A_{US}$, against log GDP per capita in the corresponding year.

4. Also interesting is the pattern that the empirical fit of the neoclassical growth model seems to deteriorate over time.
Challenges to Callibration I

- In addition to the standard assumptions of competitive factor markets, we had to assume:
  - no human capital externalities, a Cobb-Douglas production function, and a range of approximations to measure cross-country differences in the stocks of physical and human capital.

- The calibration approach is in fact a close cousin of the growth-accounting exercise (sometimes referred to as “levels accounting”).

- Imagine that the production function that applies to all countries in the world is

\[ F(K_j, H_j, A_j), \]

- Assume countries differ according to their physical and human capital as well as technology—but not according to \( F \).
Challenges to Callibration II

- Rank countries in descending order according to their physical capital to human capital ratios, \( K_j / H_j \). Then

\[
\hat{x}_{j,j+1} = g_{j,j+1} - \bar{\alpha}_{K,j,j+1} g_{K,j,j+1} - \bar{\alpha}_{L,j,j+1} g_{H,j,j+1},
\]

where:

- \( g_{j,j+1} \): proportional difference in output between countries \( j \) and \( j + 1 \),
- \( g_{K,j,j+1} \): proportional difference in capital stock between these countries and
- \( g_{H,j,j+1} \): proportional difference in human capital stocks.
- \( \bar{\alpha}_{K,j,j+1} \) and \( \bar{\alpha}_{L,j,j+1} \): average capital and labor shares between the two countries.

- The estimate \( \hat{x}_{j,j+1} \) is then the proportional TFP difference between the two countries.
Levels-accounting faces two challenges.

1. Data on capital and labor shares across countries are not widely available. Almost all exercises use the Cobb-Douglas approach (i.e., a constant value of $\alpha_K$ equal to $1/3$).

2. The differences in factor proportions, e.g., differences in $K_j/H_j$, across countries are large. An equation like (16) is a good approximation when we consider small (infinitesimal) changes.
In this lecture, the focus has been on proximate causes—importance of human capital, physical capital and technology.

Let us now return to the list of potential fundamental causes discussed in the first lecture:

1. luck (or multiple equilibria)
2. geographic differences
3. institutional differences
4. cultural differences

Do we need to worry about the relationship between these fundamental causes and the correlates of growth? In what way? Where is theory useful?
Conclusions

- Message is somewhat mixed.
  - On the positive side, despite its simplicity, the Solow model has enough substance that we can take it to data in various different forms, including TFP accounting, regression analysis and calibration.
  - On the negative side, however, no single approach is entirely convincing.
- Complete agreement is not possible, but safe to say that consensus favors the interpretation that cross-country differences in income per capita cannot be understood solely on the basis of differences in physical and human capital.
- Differences in TFP are not necessarily due to technology in the narrow sense.
- It is also useful and important to think about fundamental causes, what lies behind the factors taken as given either Solow model.