ONLINE APPENDIX STARTS HERE.
NOT FOR PUBLICATION.

**Appendix Table 1: Excluded RCT studies**

| Country | Program name | Article | Government program? | Data on all adults? | Other |
|---|---|---|---|---|---|
| Burkina Faso | NCTPP | Akresh, De Walque and Kazianga (2013) | Yes | Yes | Data not available. |
| Columbia | SCAE | Barrera-Osorio et al (2011) | **No** | Unclear | |
| Ecuador | BDH | Edmonds and Schady (2012), Schady and Caridad Araujo (2008) | Yes | Unclear | Data not yet available. |
| Kenya | CT-OVC | Asfaw et al (2014) | Yes | Yes | Sample selected differently in control. |
| Kenya | Give Directly | Haushofer and Shapiro (2013) | **No** | Yes | |
| Malawi | SCT program | Covarrubias et al (2012) | Yes | Yes | 8 clusters |
| Malawi | | Baird, McIntosh, and Ozler (2011) | **No** | **No** | |
| Nicaragua | Atención a Crisis | Macours et al (2012) | Yes | **No** | |
| Tanzania | TASAF | Evans, Hausladen, Kosec, and Reese (2014) | Yes | Yes | Data not yet available. |
| Uganda | Cash transfer for pre-school | Gilligan and Roy (2013) | **No** | Unclear | |
| Zambia | Child Program | American Institute for Research (2013) | Yes | Yes | Data not available. |

## Appendix Table 2: Balance Check

| | Honduras PRAF | | | Morocco Tayssir | | | Philippines PPP | | |
|---|---|---|---|---|---|---|---|---|---|
| | Control | Treat | Diff | Control | Treat | Diff | Control | Treat | Diff |
| *Panel A: Demographic Characteristics of Individuals Age 16-65* | | | | | | | | | |
| Male | 0.51 | 0.50 | -0.01 | 0.48 | 0.48 | 0.00 | 0.53 | 0.52 | -0.01 |
| | | | (0.01) | | | (0.01) | | | (0.01) |
| Age | 33.47 | 33.58 | 0.11 | 34.53 | 34.49 | -0.00 | 33.07 | 32.84 | -0.27 |
| | | | (0.32) | | | (0.23) | | | (0.30) |
| Years of Education | 3.07 | 3.41 | 0.35* | 1.59 | 1.60 | 0.03 | 6.32 | 6.35 | 0.10 |
| | | | (0.19) | | | (0.10) | | | (0.20) |
| Married | | | | 0.65 | 0.65 | -0.00 | 0.60 | 0.60 | -0.00 |
| | | | | | | (0.01) | | | (0.02) |
| Divorced | | | | 0.01 | 0.00 | 0.00 | 0.01 | 0.01 | -0.00 |
| | | | | | | (0.00) | | | (0.00) |
| Widow | | | | | | | 0.03 | 0.03 | 0.01* |
| | | | | | | | | | (0.01) |
| Single | | | | 0.32 | 0.32 | 0.00 | 0.37 | 0.36 | -0.01 |
| | | | | | | (0.01) | | | (0.02) |
| # People in HH | 5.61 | 5.81 | 0.19 | 6.64 | 6.57 | -0.03 | 6.04 | 6.13 | 0.09 |
| | | | (0.13) | | | (0.09) | | | (0.13) |
| P-value, joint significance | | | 0.14 | | | 0.94 | | | 0.38 |
| *Panel B: Working Variables for Individuals Age 16-65 (Controlling for Demographic Characteristics)* | | | | | | | | | |
| Worked Last Week | | | | 0.39 | 0.39 | -0.00 | | | |
| | | | | | | (0.01) | | | |
| Hours Worked Per Week | | | | 18.39 | 18.30 | -0.07 | | | |
| | | | | | | (0.48) | | | |
| Worked For Self/Family | | | | 0.14 | 0.14 | -0.01 | | | |
| | | | | | | (0.01) | | | |
| Worked Out of HH | | | | 0.25 | 0.25 | 0.00 | | | |
| | | | | | | (0.01) | | | |
| P-value, joint significance | | | | | | 0.93 | | | |

Notes: This table reports the results of a balance check between treatment and control for each study sample. Data for PRAF and PPP come from endline; data from all other programs come from baseline. For each program, the column "Ctl" is mean in control areas, "Treat" is mean in treatment areas, and "Diff" is difference between treatment and control, controlling for strata fixed effects (for all programs but PAL and Progresa), with standard errors clustered at the randomization unit. In Panel B, "Diff" also controls for demographic variables. The p-value from a joint test of demographic (Panel A) and working (Panel B) variables for each program is provided when possible.

## Appendix Table 3: Balance Check, Continued

| | Mexico PAL | | | Indonesia PKH | | | Nicaragua RPS | | | Mexico Progresa | | |
|---|---|---|---|---|---|---|---|---|---|---|---|---|
| | Control | Treat | Diff | Control | Treat | Diff | Control | Treat | Diff | Control | Treat | Diff |
| *Panel A: Demographic Characteristics of Individuals Age 16-65* | | | | | | | | | | | | |
| Male | 0.47 | 0.48 | 0.01 | 0.50 | 0.50 | 0.00 | 0.52 | 0.52 | 0.00 | 0.49 | 0.49 | 0.01* |
| | | | (0.01) | | | (0.00) | | | (0.01) | | | (0.00) |
| Age | 34.51 | 35.02 | 0.51 | 35.24 | 35.20 | -0.04 | 32.07 | 31.91 | -0.14 | 34.04 | 33.87 | -0.17 |
| | | | (0.47) | | | (0.10) | | | (0.32) | | | (0.24) |
| Years of Education | 5.26 | 5.02 | -0.24 | 5.88 | 5.81 | -0.09 | 2.34 | 2.18 | -0.25 | 3.66 | 3.65 | -0.01 |
| | | | (0.29) | | | (0.07) | | | (0.19) | | | (0.13) |
| Married | 0.69 | 0.68 | -0.00 | 0.76 | 0.76 | 0.00 | 0.61 | 0.63 | 0.02 | 0.69 | 0.68 | -0.00 |
| | | | (0.02) | | | (0.01) | | | (0.02) | | | (0.01) |
| Divorced | 0.04 | 0.03 | -0.00 | 0.01 | 0.01 | -0.00 | 0.05 | 0.06 | 0.01 | 0.02 | 0.02 | -0.00 |
| | | | (0.01) | | | (0.00) | | | (0.01) | | | (0.00) |
| Widow | 0.03 | 0.03 | -0.01 | 0.04 | 0.04 | -0.00 | 0.02 | 0.02 | -0.00 | 0.04 | 0.04 | -0.00 |
| | | | (0.00) | | | (0.00) | | | (0.00) | | | (0.00) |
| Single | 0.24 | 0.26 | 0.02 | 0.18 | 0.19 | 0.00 | 0.31 | 0.29 | -0.03* | 0.25 | 0.26 | 0.01 |
| | | | (0.01) | | | (0.01) | | | (0.01) | | | (0.01) |
| # People in HH | 4.74 | 4.64 | -0.10 | 5.16 | 5.12 | -0.05 | 6.22 | 6.05 | -0.14 | 5.61 | 5.60 | -0.01 |
| | | | (0.18) | | | (0.04) | | | (0.14) | | | (0.07) |
| P-value, joint significance | | | 0.33 | | | 0.41 | | | 0.58 | | | 0.24 |
| *Panel B: Working Variables for Individuals Age 16-65 (Controlling for Demographic Characteristics)* | | | | | | | | | | | | |
| Worked Last Week | 0.51 | 0.51 | -0.00 | 0.59 | 0.59 | -0.00 | 0.58 | 0.58 | -0.00 | 0.51 | 0.53 | 0.02** |
| | | | (0.02) | | | (0.01) | | | (0.02) | | | (0.01) |
| Hours Worked Per Week | 22.84 | 21.12 | -1.85* | | | | 23.81 | 23.49 | -0.43 | 22.34 | 22.96 | 0.44 |
| | | | (0.97) | | | | | | (0.84) | | | (0.49) |
| Worked For Self/Family | 0.23 | 0.25 | 0.02 | | | | 0.27 | 0.25 | -0.02 | 0.10 | 0.14 | 0.04*** |
| | | | (0.03) | | | | | | (0.02) | | | (0.01) |
| Worked Out of HH | 0.28 | 0.26 | -0.02 | | | | 0.31 | 0.33 | 0.02 | 0.38 | 0.35 | -0.03** |
| | | | (0.03) | | | | | | (0.03) | | | (0.01) |
| P-value, joint significance | | | 0.04 | | | | | | 0.62 | | | 0.01 |

Notes: This table reports the results of a balance check between treatment and control for each study sample. Data for PRAF and PPP come from endline; data from all other programs come from baseline. For each program, the column "Control" is mean in control areas, "Treat" is mean in treatment areas, and "Diff" is difference between treatment and control, controlling for strata fixed effects (for all programs but PAL and Progresa), with standard errors clustered at the randomization unit. In Panel B, "Diff" also controls for demographic variables. The p-value from a joint test of demographic (Panel A) and working (Panel B) variables for each program is provided when possible.

## Appendix Table 4:

## Experimental Estimates of the Impact of Cash Transfer Programs on Household and Private Market Work, by Gender

| | Honduras PRAF | Morocco Tayssir | Philippines PPPP | Mexico PAL | Indonesia PKH | Nicaragua RPS | Mexico Progresa |
|---|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| | | | *Panel A. Worked in household - MEN* | | | | |
| Treatment Effect | 0.0632* | -0.0177 | -0.0034 | 0.0109 | - | 0.0559** | -0.0303** |
| | (0.0331) | (0.0279) | (0.0246) | (0.0329) | - | (0.0250) | (0.0139) |
| Observations | 4,276 | 13,879 | 2,377 | 7,306 | - | 6,632 | 89,423 |
| Control Group Mean | 0.67 | 0.63 | 0.31 | 0.30 | - | 0.46 | 0.10 |
| | | | *Panel B. Worked in household – WOMEN* | | | | |
| Treatment Effect | -0.0076 | 0.0006 | -0.0379* | 0.0020 | - | -0.0008 | -0.0173* |
| | (0.0186) | (0.0360) | (0.0195) | (0.0159) | - | (0.0125) | (0.0102) |
| Observations | 4,207 | 15,951 | 2,150 | 8,292 | - | 6,347 | 92,895 |
| Control Group Mean | 0.16 | 0.42 | 0.19 | 0.06 | - | 0.05 | 0.03 |
| | | | *Panel C. Worked outside the household - MEN* | | | | |
| Treatment Effect | -0.0156 | 0.0169 | 0.0367 | -0.0105 | - | -0.0584* | 0.0378** |
| | (0.0327) | (0.0272) | (0.0269) | (0.0333) | - | (0.0321) | (0.0190) |
| Observations | 4,279 | 13,879 | 2,377 | 7,306 | - | 6,632 | 89,423 |
| Control Group Mean | 0.38 | 0.32 | 0.39 | 0.46 | - | 0.47 | 0.70 |
| | | | *Panel D. Worked outside the household - WOMEN* | | | | |
| Treatment Effect | -0.0458* | -0.0161** | 0.0256 | -0.0091 | - | -0.0268 | -0.0008 |
| | (0.0233) | (0.0064) | (0.0198) | (0.0133) | - | (0.0207) | (0.0059) |
| Observations | 4,207 | 15,951 | 2,150 | 8,292 | - | 6,347 | 92,895 |
| Control Group Mean | 0.13 | 0.02 | 0.18 | 0.10 | - | 0.11 | 0.08 |
| Method | endline | DD | endline | DD | DD | DD | DD |

Notes: This table replicates Table 4, separating results by gender. See Table 3 for specification details. *** p<0.01, ** p<0.05, * p<0.1

## Appendix Table 5: Pooled Impact of Cash Transfer Programs on Household and Private Market Work Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
|---|---|---|---|---|---|---|---|
| | **Weighted control mean** | **Effect size ($\tau_p$)** | | | **Elasticity ($\tau_p$)** | | |
| *Statistic of the posterior distribution:* | | Mean | 5th percentile | 95th percentile | Mean | 5th percentile | 95th percentile |
| | | | ***Panel A. Worked in household*** | | | | |
| Full sample: | 0.28 | 0.000 | -0.055 | 0.055 | 0.00 | -1.43 | 1.42 |
| For Men: | 0.41 | 0.011 | -0.085 | 0.105 | 0.19 | -1.50 | 1.86 |
| For Women: | 0.15 | -0.011 | -0.047 | 0.024 | -0.56 | -2.28 | 1.18 |
| | | | ***Panel B. Worked outside the household*** | | | | |
| Full sample: | 0.27 | -0.004 | -0.070 | 0.056 | -0.11 | -1.87 | 1.50 |
| For Men: | 0.45 | 0.005 | -0.081 | 0.091 | 0.08 | -1.31 | 1.48 |
| For Women: | 0.10 | -0.014 | -0.067 | 0.043 | -1.00 | -4.72 | 3.03 |

Table Notes. This table reports results from a Bayesian hierarchical model used to aggregate the results from the seven programs. The impact for each program from Table 3 or Table 5 is first scaled according to the size of the transfer, such that for each program the scaled coefficient corresponds to a transfer worth 13.6% of consumption. (The program transfer size is defined as the average transfer value relative to average consumption.) Column (1) reports the mean of the row variable in the control group at endline, averaged over the seven programs. Columns (2)-(4) present the mean, and the 5th and 95th percentiles of the posterior distribution of the site effect $\tau_p$, which measures the impact for a hypothetical new program. Columns (5)-(7) report the same statistics for the elasticity of the work outcome with respect to the size of the cash transfer. Bayesian posteriors are computed using the rstan package, 20,000 iterations on 4 chains, thinning the result by a factor of two.

**Appendix Figure 1: Experimental Estimates of Cash Transfers on Work Outcomes (conditional on work)**

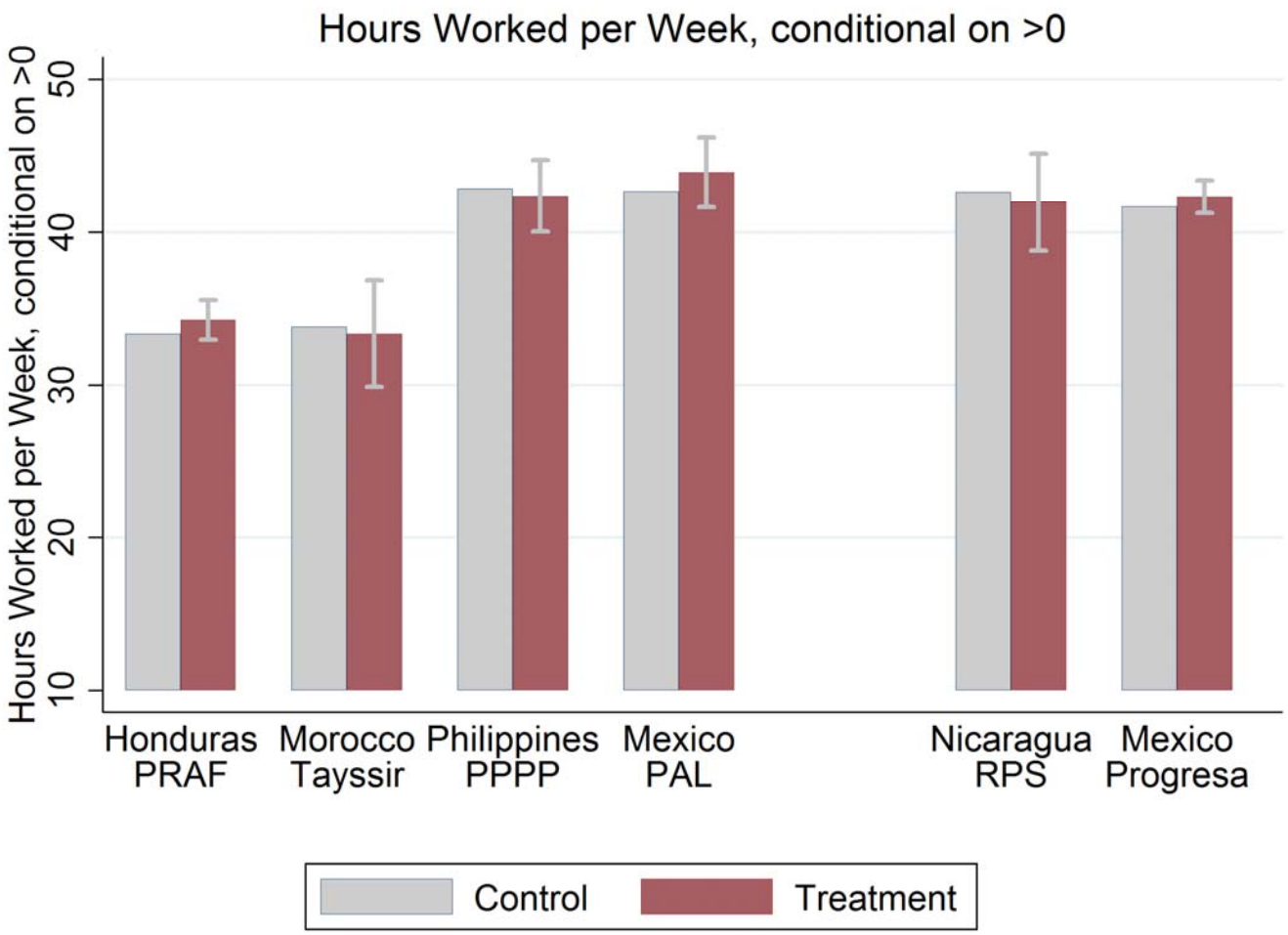


**Figure Notes.** The "Control" (gray) bars report the mean of the outcome variable (probability of work and hours worked in Panels A and B, respectively) in the control group, at endline. The "Treatment" (dark red) bars report the control mean plus the treatment effect from a regression similar to that in Table 3 Panel B, for the conditional outcome variable. The gray segments represent 95% confidence intervals.

## Appendix 1: Details of Pooling Estimation

Following Rubin (1981), we fit a hierarchical model on the estimated treatment effects $\hat{\tau}_p$ obtained from equation (2). Under the standard asymptotic normality approximation for $\hat{\tau}_p$, the parent distribution and the estimated parameters are linked by the following system of equations:

$$\tau_p \sim N(\tau, \sigma_\tau)$$

$$\hat{\tau}_p \sim N(\tau_p, \widehat{se}_p)$$

where $\widehat{se}_p$ is the estimated standard error corresponding to $\hat{\tau}_p$. Following Meager (2016), we use a Bayesian approach to estimate $\tau$ and $\sigma_\tau$. We assume flat, uninformative priors on these parameters, and obtain the posterior joint distribution given data using a Markov Chain Monte Carlo (MCMC) algorithm.[32] Intuitively, this method creates a sequence whose stationary distribution is exactly the joint posterior of $(\tau, \sigma_\tau)$. We focus on reporting statistics of the posterior distribution of $\tau_p$, which represents the impact in a hypothetical new cash transfer program. Note that $\tau_p$ includes the heterogeneity in program effects quantified by $\sigma_\tau$. We report the mean, 5th and 95th percentiles of the marginal distribution of $\tau_p$. The latter quantities $\tau_p^5$ and $\tau_p^{95}$ are defined implicitly by:

$$\Pr_{\tau,\sigma_\tau}\left(\tau_p < \boldsymbol{\tau_p^{perc}} \mid \tau_p \sim N(\tau, \sigma_\tau)\right) = \int_{-\infty}^{\infty} \Phi\left(\left(\boldsymbol{\tau_p^{perc}} - \tau\right)/\sigma_\tau\right) d\mu(\tau, \sigma_\tau) = \frac{perc}{100} \text{ for } perc = 5, 95$$

where $\Phi(\cdot)$ is the standard normal cdf and $\mu(\cdot,\cdot)$ is the estimated posterior distribution. For any $\tau$ and $\sigma_\tau$, the probability that the treatment effect is below $\tau_p^{perc}$ is given by $\Phi\left(\left(\tau_p^{perc} - \tau\right)/\sigma_\tau\right)$. We then integrate this function over the entire posterior to obtain the posterior probability.

[32] In practice, we use the RStan package, which implements a Hamiltonian MCMC with No-U-Turn (NUTS) sampling. We run 4 chains with 20,000 iterations and a thinning factor equal to 2.

## Appendix 2: Program Notes

*Program Notes for Honduras PRAF*

**Data Source.** Endline survey data from Glewwe and Olinto (2004); treatment and randomization strata from Galiani and McEwan (2013).

**Program Eligibility.** The 70 poorest municipalities in Honduras were eligible for PRAF[33]. Within eligible municipalities, households with pregnant women or children younger than three were eligible for the health transfer, and households with children 6-12 enrolled in grade 1-4 were eligible for the education transfer.

**Original Study Randomization.** 70 municipalities were grouped into five strata based on their mean height-for-age z-score for first graders. Within each stratum, four municipalities were randomly assigned to receive CCTs, four to receive CCTs plus direct investments in health and education, two to receive direct investments only, and four to control. Randomization occurred publically.

**Original Study Sample.** Within each municipality, 8 communities were randomly selected, and 10 dwellings (which could contain multiple households) were randomly selected to be interviewed at baseline. The endline sample consists of 92% of the 5,748 households from baseline. Additionally, household members who left their baseline household were followed if they were pregnant women, lactating mothers, or children age 0-16.

**Sample restriction**. We restrict to municipalities assigned to CCTs only or control. We restrict to adults between 16 and 65 at the time of the survey. Individuals defined as the household's domestic workers are dropped from the sample (57 observations).

**Variable Definitions**. The household roster asks whether each member worked in the last week for a private job and if each member worked in the last week for self/family. Household members are also asked for the work activity to which they dedicated the most time last week, the days spent working on this activity, and the hours per day spent working on this activity.

We code the "Worked last week" dummy is equal to 1 if the household member reports working in a private job, working for self/family, or working 1-7 days on a work activity. The "Worked for Self/Family" dummy is equal to 1 if the household member reports working for self/family, and the "Worked Out of HH" dummy is equal to 1 if the household member reports working for a private job. "Hours per week" is obtained by multiplying hours per day and days per week, censoring at 98 hours per week, and filling with zero if the individual did not work.

The labor force dummy is equal to 1 if the person has ever worked.

We code age, gender, and years of education, as well as household size, from the household roster.

**Other notes**. We do not use baseline data from PRAF because treatment and control areas were interviewed in different coffee growing seasons, which influences labor.

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[33] 298 municipalities were sorted based on mean height-for-age z-score for first graders; eligible municipalities had a z-score below the cutoff od -2.304. Three municipalities were excluded because of distance and cost, leaving a final sample of 70 municipalities.

*Program Notes for Morocco Tayssir*

**Data Source.** Data from Benhassine, Devoto, Duflo, Dupas, and Pouliquen (2015).

**Program Eligibility.** All households in treatment school sectors were eligible.

**Original Study Randomization.** 314 school sectors were randomized into control (59 sectors) and treatment (255), stratified by 46 regions.

**Original Study Sample.** In each school cluster, 6 households were randomly selected from a list of households in the school's vicinity that had at least one child enrolled in school, and 2 households were randomly selected from a list of households with no child currently enrolled in school but at least one child of school-age who had enrolled at some point but dropped out within the previous three years.

**Sample restriction**. We restrict to adults between 16 and 65 at the time of the survey. No domestic workers appear in the original sample.

**Variable Definitions**. There are two types of questions on work activities. The household roster asks the main activity of each member during the past 30 days, with the following options: (1) worked throughout the period as self-employed, (2) was employed throughout the period, (3) worked occasionally, (4) was looking for work, (5) domestic tasks, (6) studied, (7) was sick, (8) is retired, (9) did not work, (10) other.

Another section (G) asks for each member about the top three paid and unpaid work activities during the past 30 days. For each activity, it contains questions on the occupation, the work relation ((1) casual work, (2) permanent employee, (3) paid apprentice, (4) unpaid apprentice, (5) self-employed, (6) sharecropping, (7) unpaid help in household activity, (8) paid help in household activity, other), the number of days in the past 30 days, and the average number of hours per working day in this activity.

Section G has more detailed questions and thus it is in principle preferable to use it to code the work variables. In practice, we use the roster (main activity) question at baseline, and the questions in section G at endline. The reason is that in the baseline survey, the occupation question in section G also has an option "domestic tasks"; this option disappears in the endline. In practice, this makes it difficult to distinguish in the baseline survey, for women, between bona fide domestic tasks and agriculture and ranching. [34] Our results are robust to using the roster question for both baseline and endline (available upon request).

At baseline, we code the "Worked last week" dummy equal to 1 for options 1, 2 and 3 of the roster question (main activity). The "Worked for Self/Family" dummy is equal to 1 for options 1, and the "Worked Out of HH" dummy is equal to 1 for options 2 and 3. At endline, we code the "Worked last week" dummy equal to 1 as long as section G contains at least one work activity. The "Worked for

[34] At baseline, 88% of adult women reported at least one activity in section G – this includes domestic tasks. By comparison, 44% of adult women reported a work activity in section G at endline, when domestic tasks was no longer an option. For other programs we do not code domestic tasks as work, hence at baseline we decided to use only the roster (main reported) question. This leads to classifying 3% of women as working at baseline, and 44% at endline. For men, these numbers are 78% and 85%.

Self/Family" dummy is equal to 1 if at least one of the three activities has a work relation 5, 6, 7 or 8, and the "Worked Out of HH" dummy is equal to 1 if at least one of the three activities has a work relation 1, 2, 3 or 4. Hours per month are calculated as the total over the three work activities in section G; at baseline, hours per month are set to zero if the "Worked last week" dummy is equal to 0.. "Hours per week" is obtained by multiplying by 7/30 and censoring at 98 hours per week.

The labor force dummy is equal to 1 if the person worked in the previous month, or if they searched for work (according to the main activity roster questions).

We code age, gender, years of education and marital status, as well as household size, from the household roster.

**Other notes**. Years of education is asked for persons above 15 years and 17 years at baseline and endline, respectively. Thus, we are missing this data for 16 year olds at endline.

*Program Notes for Philippines PPPP*

**Data Source.** Data from Chaudhury, Friedman and Onishi (2013).

**Program Eligibility.** Beneficiaries were selected using a combination of geographical targeting and proxy means testing (PMT), using data from the National Household Targeting System for Poverty Reduction (NHTS-PR). Households are eligible if they have a pregnant mother at the time of the Household Assessment by NHTS-PR and/or children between 0-14 years of age.

**Original Study Randomization.** 130 villages (barangay) were randomized into 65 treatment and 65 control villages, stratified by 8 municipalities.

**Original Study Sample.** The original study included 4 sample groups.

**Sample restriction**. We first restrict attention to sample group 1, namely 10 random eligible households in each study village. We restrict to adults between 16 and 65 at the time of the survey.

**Variable Definitions**. The survey contains a question on the whether the respondent did any work or business at least one hour during the past 7 days. This is our "Worked last week" dummy. A follow-up question asks the sector of work, with options: (0) worked for private household, (1) worked for private establishment, (2) work for government corporation, (3) self-employed without any paid employee, (4) employer in own family-operated farm or business, (5) worked with pay on own family-operated farm or business, (6) worked without pay on own family-operated farm or business, (8) don't know, (9) no response. The "Worked for Self/Family" dummy is equal to 1 for options 3, 4, 5 and 6, and the "Worked Out of HH" dummy is equal to 1 for options 0, 1 and 2. We use a follow-up question on the total number of hours from all jobs during the past 7 days to calculate the "Hours per week" variable, which we top code at 98 hours per week. The labor force dummy is equal to 1 if the person worked or was unemployed (searched or waited for job) in the previous week.

We code age, gender, years of education and marital status, as well as household size, from the household roster.

*Program Notes for Mexico PAL*

**Data Source.** Data from Cunha (2014).

**Program Eligibility.** Poor households in eligible localities[35] (around 89% of sample population); however, household targeting was not implemented, implying all households in treated localities received PAL.

**Original Study Randomization.** 208 localities in eight states in southern Mexico were randomized into one of three treatments (25% of localities each to in-kind transfer treatment, in-kind transfer plus education treatment, or cash transfer plus education treatment[36]) or control (25% of localities). Two localities could not be surveyed due to violence.

**Original Study Sample.** Within each locality in the sample, 33 households were randomly selected to be interviewed as a panel, resulting in a baseline sample of 6,696 households.

**Sample restriction**. We restrict to localities assigned to the cash transfer treatment or control. Following Cunha (2014), we drop the one locality in our sample that received Oportunidades (rendering it ineligible for PAL).

We restrict to adults between 16 and 65 at the time of the survey. Individuals defined as the household's domestic workers (or the family of domestic workers) are dropped from the sample (2 observations).

**Variable Definitions**. The household roster asks whether each member worked last week, with the following options: (1) worked, (2) unemployed, (3) had a job but did not work, (4) looked for work, (5) was a student, (6) is devoted to care of home, (7) retired, (8) permanently unable to work.

If option 1 is chosen, a question about the type of work last week is asked, with the following options: (1) works on land or for family business without pay, (2) works on land or for family business with pay, (3) works for non-family business, (4) works for government, (5) craftsman or laborer, (6) mason, (7) day laborer, (8) foreman, (9) domestic employee, (10) driver, (11) street vendor, (12) own business, (13) self-employed, (14) ironing, washing, or sewing outside household, (15) other.

If option 1-7 is chosen, a follow-up question is asked to capture activities in the last week that may not be considered working, with the following options: (1) sold a product, (2) worked for a family business, (3) made products to sell, (4) washed, ironed, or sewed in exchange for pay, (5) helped with agricultural activities or raising animals, (6) another type of work for pay.

Finally, a question is asked about total hours worked last week, including both the main job and activites mentioned in the follow-up question. This is censored at 98 hours per week.

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[35] Villages are eligible for PAL if they have fewer than 2,500 inhabitants, are highly marginalized (classified by Census Bureau), and do not receive aid from Liconsa (subsidized milk program) or Oportunidades (CCT that originated as Progresa). PAL villages tend to be poorer and more rural than Progresa villages.

[36] The education portion of treatment was the provision of health, hygiene and nutrition classes; however, attendance was not required to receive benefits, and Cunha (2014) finds no difference in consumption between households in the cash treatment that attend and do not attend classes.

We code the "Worked last week" dummy is equal to 1 for option 1 of the first question (worked last week) and any of the options for the third question (activities last week). The "Worked for Self/Family" dummy is equal to 1 for options 1, 2, 11, 12, and 13 for the second question (type of work), and the "Worked Out of HH" dummy is equal to 1 for options 3-10 and 14. "Hours per week" is taken directly from the question about hours, filling with zero if the individual did not work. The labor force dummy is equal to 1 if the person worked, had a job but did not work, or was unemployed in the previous week.

We code age, gender, years of education and marital status, as well as household size, from the household roster.

**Other notes**. While the PAL evaluation was designed as a panel, we do not restrict to individuals or households surveyed at both baseline and endline.

*Program Notes for Indonesia PKH*

**Data Source.** Data from World Bank Office Jakarta (2011).

**Program Eligibility.** PKH eligibility was determined with a PMT test. All surveyed households were eligible.

**Original Study Randomization.** 360 sub-districts (kecamatan) were randomized into 180 treatment and 180 control, stratified by 44 districts (kabupaten).

**Original Study Sample.** Statistics Indonesia surveyed poor and extremely poor households, which were drawn from the 2005 BLT beneficiaries list (known as PPLS05).[37] All BLT households that did not meet the PKH criteria were dropped. Additional poor households were added to the PPLS05 list; this resulted in approximately an additional 5 percent households.

**Sample restriction**. We restrict to adults between 16 and 65 at the time of the survey, and exclude domestic workers (30 individuals).

**Variable Definitions**. The survey only contains a question on the main activity performed last week, with options: (1) Employed, (2) Attend school, (3) Look after HH, (4) Retired, (5) Unemployed, (6) Under five, and (7) Other. We code the "Worked last week" dummy is equal to 1 for option 1. We cannot code any of the other work variables. The labor force dummy is equal to 1 if the person worked or was unemployed in the previous week.

We code age, gender, years of education and marital status, as well as household size, from the household roster.

[37] Bantuan Langsung Tunai (BLT) is an unconditional cash transfer program.

**Data Source.** Data from International Food Policy Research Institute (2012) <http://hdl.handle.net/1902.1/17535>

**Program Eligibility.** Within eligible comarcas[38], nearly all households were eligible for the food security transfer[39], and households with children age 7-13 who had not completed fourth grade were eligible for the education transfer.

**Original Study Randomization.** The 42 comarcas were sorted into seven strata based on their marginality index scores. Within each stratum, three comarcas were randomly assigned to treatment and three were randomly assigned to control. Randomization occurred publically.

**Original Study Sample.** A stratified random sample was selected from all 42 comarcas. Within each comarca, 42 households were randomly selected using a pre-survey census carried out three months prior. 1,581 (90% of sample) households were surveyed at baseline; followup surveys were restricted to the sample interviewed at baseline.

**Sample restriction**. We restrict to adults between 16 and 65 at the time of the survey who are classified as household members. Individuals defined as the household's domestic workers are dropped from the sample (2 observations).

**Variable Definitions**.

The household roster asks whether each member worked last week, with the following options: (1) worked, (2) had a job but did not work, (3) worked for a business, property or ranch without pay, (4) did not work. If option 4 is chosen, a follow-up question is asked to capture activities in the last week that may not be considered working, with the following options: (1) sold a product, (2) worked for a business, property or ranch without pay, (3) made products to sell, (4) washed, ironed, or cooked in exchange for pay, (5) helped with agricultural activities or raising animals.

Household members are also asked for their position in their **main** job, with the following options: (1) rural laborer or field worker, (2) non-agricultural worker, (3) self-employed, (4) business owner, (5) unpaid family worker, (6) unpaid non-family worker, (7) cooperative member, (8) ejidatario[40], (9) other. Finally, they are asked for days worked last week in their main job, as well as average hours worked per pay.

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[38] Six municipalities (containing 59 rural comarcas total) in rural areas of all 17 departments of Nicaragua were selected because they had a small-scale participatory development program. A marginality index was constructed for rural comarcas in these areas based on average family size, percent without piped water, percent without latrine, and percent of people over 5 who are illiterate. The 42 comarcas with the highest marginality score (i.e. most impoverished) were eligible for RPS.

[39] Less than 3% of households were excluded ex ante because they owned a vehicle and/or owned more than 14.1 hectares of land. None of these households were sampled. Additionally, less than 4% of households were excluded after program registration because (1) contained a single, non-disabled man or woman, (2) had significant economic resources or a business, and/or (3) omitted or falsified information during the RPS population census.

[40] *Ejidos* are communal land ownership schemes in which community members individually possess a specific parcel of land.

We code the "Worked last week" dummy is equal to 1 for options 1 and 3 of the first question (worked last week) and options 1 through 5 of the second question (activities last week). The "Worked for Self/Family" dummy is equal to 1 for options 3, 4, and 5 for the third question (position in main job), and the "Worked Out of HH" dummy is equal to 1 for options 1 and 2 for the third question, or options 1, 3, or 4 of the second question (activities last week). "Hours per week" is obtained by multiplying hours per day and days per week, censoring at 98 hours per week, and filling with zero if the individual did not work. The labor force dummy is equal to 1 if in the previous week the person worked, was employed but did not work due to holiday or sickness, or was unemployed, including waiting for the agricultural season to start, waiting for a job offer, or waiting to begin a new job,.

We code age, gender, years of education[41], and marital status, as well as household size, from the household roster.

**Other notes**. 37 households living in control comarcas that appearred to have been program beneficiaries were not included in the data. While the RPS evaluation was designed as a panel, we do not restrict to individuals or households surveyed in every round.

*Program Notes for Mexico Progresa*

**Data Source.** Data from Secretary of Social Development (SEDESOL), Government of Mexico <http://evaluacion.oportunidades.gob.mx:8010>

**Program Eligibility.** Poor households[42] in eligible localities[43], with children enrolled in grades 3-9.

**Original Study Randomization.** 506 localities in seven states were randomized into treatment (320 localities) or control (186 localities).

**Original Study Sample.** All households (24,077 at baseline) in treatment and control localities, including both eligibles and ineligibles, were interviewed every six months. We use data from November 1997 (pre-program), November 1998, June 1999, and November 1999. The pre-program March 1998 survey is excluded because it only collects individual-level data for children 6-18.

[41] The pre-baseline census collected education variables for original household members, while the surveys collected education variables for new household members. We use survey data if available, and fill with census data if not.

[42] Initially, definition of poor included 52% of households; this was revised to include 78% of households before treatment started. Due to administrative errors and delays, around 65% of households actually received transfers within the evaluation timeframe (Skoufias and Di Maro, 2006). We use the broader definition of eligibility in our analysis.

[43] Eligibility was based on a marginality index consisting of share of illiterate adults, share of population working in the primary sector, average occupants per room, and share of dwellings without water, without drainage systems, without electricity, and with dirt floor. Geographical location, distance between localities, and the existence of health and school infrastructure were also considered (Skoufias, Davis, and Behrman, 1999).

**Sample restriction**. We restrict to adults between 16 and 65 at the time of the survey. Individuals defined as the household's domestic workers are dropped from the sample (91 observations across 4 survey rounds).

**Variable Definitions**. The household roster asks whether each member worked last week, with the following options: (1) worked, (2) had a job but did not work, (3) worked for a business, property or ranch without pay, (4) did not work. If option 4 is chosen, a follow-up question is asked to capture activities in the last week that may not be considered working, with the following options: (1) sold a product, (2) worked for a business, property or ranch without pay, (3) made products to sell, (4) washed, ironed, or cooked in exchange for pay, (5) helped with agricultural activities or raising animals.

Household members are also asked for their position in their **main** job, with the following options: (1) rural laborer or field worker, (2) non-agricultural worker, (3) self-employed, (4) business owner, (5) unpaid family worker, (6) unpaid non-family worker, (7) cooperative member, (8) ejidatario[44], (9) other. Finally, they are asked for days worked last week in their main job, as well as average hours worked per pay.

We code the "Worked last week" dummy is equal to 1 for options 1 and 3 of the first question (worked last week) and options 1 through 5 of the second question (activities last week). The "Worked for Self/Family" dummy is equal to 1 for options 3, 4, and 5 for the third question (position in main job), and the "Worked Out of HH" dummy is equal to 1 for options 1 and 2 for the third question, or options 1, 3, or 4 of the second question (activities last week). "Hours per week" is obtained by multiplying hours per day and days per week, censoring at 98 hours per week[45], and filling with zero if the individual did not work. The labor force dummy is equal to 1 if the person worked or had a job but did not work in the previous week.

We code age, gender, years of education[46] and marital status, as well as household size, from the household roster.

**Other notes**. While the Progresa evaluation was designed as a panel, we do not restrict to individuals or households surveyed in every round.

[44] *Ejidos* are communal land ownership schemes in which community members individually possess a specific parcel of land.
[45] June 1999 does not ask for hours, and so is excluded from this portion of analysis.
[46] Education is asked about for all members in baseline, and for new household members and/or individuals 6-18 in followups. We match individuals across surveys and use their earliest available education variable for all waves.